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Excerpt

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CHAPTER I

PRELIMINARIES

1. Rearrangement invariant spaces

Let (Ω, \mathcal{F}, P) be a probability space, i.e. a measure space such that $P(\Omega) = 1$. A measurable function X on Ω is called a *random variable* (r.v.). We write $X \stackrel{d}{=} Y$ if these r.v.s are identically distributed. This means that $P\{X < x\} = P\{Y < x\}$ for every $x \in \mathbf{R}$. The *mean value* of a r.v. X is defined by the formula

$$EX = \int_{\Omega} X(\omega) dP(\omega).$$

The *distribution* of a r.v. X is the measure F on \mathbf{R} defined by the formula $F(h) = P\{X \in h\}$ for every measurable set $h \subset \mathbf{R}$. We write $X \in \mathcal{L}(F)$ if X has the distribution F . The function

$$f(t) = \int_{-\infty}^{\infty} \exp(itx) dF(x)$$

is called the *characteristic function* of the r.v. X .

Definition 1. A Banach space \mathbf{E} of random variables defined on (Ω, \mathcal{F}, P) is said to be *rearrangement invariant* (r.i.) if the following conditions hold:

- (i) if $|X| \leq |Y|$ and $Y \in \mathbf{E}$, then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq \|Y\|_{\mathbf{E}}$;
- (ii) if $X \stackrel{d}{=} Y$ and $Y \in \mathbf{E}$, then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} = \|Y\|_{\mathbf{E}}$.

The main information on r.i. spaces is contained in the book [29]. The spaces $L_p(\Omega)$ ($1 \leq p \leq \infty$) are rearrangement invariant. The Lorentz spaces $L_{p,q}(\Omega)$ form a wider class of r.i. spaces which consist of all r.v.s X such that

$$\|X\|_{p,q}^* = \left(\int_0^{\infty} (P\{|X| \geq x\})^{q/p} dx^q \right)^{1/q} < \infty \quad (1)$$

if $1 \leq p, q < \infty$, and,

$$\|X\|_{p,\infty}^* = \sup \left\{ x(P\{|X| \geq x\})^{1/p} : x > 0 \right\} < \infty. \quad (2)$$

The functionals (1) and (2) are not norms in general, but they are equivalent to some norms (see [52]). We have $L_{p,p} = L_p$.

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Orlicz spaces form another class of r.i. spaces. Let $N(x)$ be a convex even function on \mathbf{R} , $N(0) = 0$. The Orlicz space $L_N(\Omega)$ consists of all r.v.s X such that $EN(\lambda^{-1}X) < \infty$ for some $\lambda > 0$. The norm is defined by the formula (see [28])

$$\|X\|_{L_N} = \inf \{ \lambda > 0 : EN(\lambda^{-1}X) < 1 \}. \tag{3}$$

In the sequel the probability space (Ω, \mathcal{F}, P) is assumed to be non-atomic. The indicator of a set h is denoted by I_h .

The following statements are proved in [29].

Proposition 1. For every r.i. space \mathbf{E}

$$L_\infty(\Omega) \subset \mathbf{E} \subset L_1(\Omega)$$

and for each r.v. $X \in L_\infty(\Omega)$

$$\|X\|_{L_1(\Omega)} \leq \frac{\|X\|_{\mathbf{E}}}{\|I_\Omega\|_{\mathbf{E}}} \leq \|X\|_{L_\infty(\Omega)}.$$

Proposition 2. Let $Y \in \mathbf{E}$ and $P\{|X| \geq x\} \leq CP\{|Y| \geq x\}$ for all $x > 0$, where C is a constant. Then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq \max\{1, C\}\|Y\|_{\mathbf{E}}$.

The dual or associated space \mathbf{E}' of the r.i. space \mathbf{E} is the set of all r.v.s Y such that

$$\|Y\|_{\mathbf{E}'} \stackrel{def}{=} \sup \{ EXY : X \in \mathbf{E}, \|X\|_{\mathbf{E}} \leq 1 \} < \infty.$$

It is well known (see [29]) that \mathbf{E}' is a r.i. space. Write $\mathbf{E}'' = (\mathbf{E}')'$. We have $\mathbf{E} \subset \mathbf{E}''$ and $\|X\|_{\mathbf{E}''} \leq \|X\|_{\mathbf{E}}$ for every $X \in \mathbf{E}$. If $X \in L_\infty(\Omega)$, then $\|X\|_{\mathbf{E}''} = \|X\|_{\mathbf{E}}$. A r.i. space \mathbf{E} is said to be maximal, if $\mathbf{E}'' = \mathbf{E}$.

The decreasing rearrangement of a r.v. X is the function on $(0, 1)$ defined by the formula

$$X^*(t) = \inf \{ s > 0 : P\{|X| \geq s\} < t \}.$$

We write $X \prec Y$ for $X, Y \in L_1(\Omega)$ if for each $0 \leq t \leq 1$

$$\int_0^t X^*(s)ds \leq \int_0^t Y^*(s)ds.$$

The following assertions are proved in [29].

Proposition 3. Let a r.i. space \mathbf{E} be maximal or separable, $Y \in \mathbf{E}$ and $X \prec Y$. Then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq \|Y\|_{\mathbf{E}}$.

Proposition 4 (Calderon–Mityagin’s Theorem). Suppose that a r.i. space \mathbf{E} has the following property: if $Y \in \mathbf{E}$ and $X \prec Y$ then $X \in \mathbf{E}$ and

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$\|X\|_{\mathbf{E}} \leq \|Y\|_{\mathbf{E}}$. Let a linear operator T be bounded in $L_1(\Omega)$ and $L_\infty(\Omega)$. Then T is bounded in \mathbf{E} and

$$\|T\|_{\mathbf{E} \rightarrow \mathbf{E}} \leq \max \{ \|T\|_{L_1(\Omega) \rightarrow L_1(\Omega)}, \|T\|_{L_\infty(\Omega) \rightarrow L_\infty(\Omega)} \}.$$

Let X_n and X be r.v.s with the distributions F_n and F respectively. The sequence X_n is said to be *weakly convergent* to X if for every continuous bounded function $g(x)$ on \mathbf{R}

$$\lim_{n \rightarrow \infty} \int_{-\infty}^{\infty} g(x) dF_n(x) = \int_{-\infty}^{\infty} g(x) dF(x).$$

Weak convergence is equivalent to each of the following conditions:

- (i) $F_n(x) \rightarrow F(x)$ for all $x \in \mathbf{R}$ on which F is continuous ;
- (ii) $f_n(t) \rightarrow f(t)$ for each $t \in \mathbf{R}$, where $f_n(t)$ and $f(t)$ are the corresponding characteristic functions (see [51]).

Proposition 5. Let \mathbf{E} be a maximal r.i. space and $X_n \in \mathbf{E}$. Suppose X_n is weakly convergent to X and

$$\sup_n \|X_n\|_{\mathbf{E}} = C < \infty.$$

Then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq C$.

Proof: Let $Y \in \mathbf{E}'$ and $\|Y\|_{\mathbf{E}'} \leq 1$. According to Proposition 1, $E|Y| < \infty$. From here $Y^* \in L_1(0, 1)$ and the function

$$\Psi_Y(t) = \int_0^t Y^*(s) ds$$

is well defined on $(0, 1)$. It is clear that

$$\int_0^1 X^*(t) Y^*(t) dt = \int_0^1 X^*(t) d\Psi_Y(t).$$

We use the formulae (see [29], Ch. 2)

$$\int_0^1 X^*(t) d\Psi_Y(t) = \int_0^\infty \Psi_Y(P\{|X| \geq x\}) dx$$

and

$$\|X\|_{\mathbf{E}''} = \sup \left\{ \int_0^1 X^*(t) Y^*(t) dt : \|Y\|_{\mathbf{E}'} \leq 1 \right\}.$$

Fatou's lemma and these equalities imply that

$$\begin{aligned} \int_0^\infty \Psi_Y(P\{|X| \geq x\})dx &\leq \sup_n \int_0^\infty \Psi_Y(P\{|X_n| \geq x\})dx \\ &= \sup_n \int_0^1 X_n^*(t)Y^*(t)dt \leq \sup_n \|X_n\|_{\mathbf{E}} = C. \end{aligned}$$

So,

$$\sup \left\{ \int_0^1 X^*(t)Y^*(t)dt : \|Y\|_{\mathbf{E}'} \leq 1 \right\} \leq C < \infty.$$

Hence $X \in \mathbf{E}'' = \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq C$. \square

If \mathbf{E} is a r.i. space, then the norm $\|I_h\|_{\mathbf{E}}$ depends on $P(h)$ only. Therefore the function

$$\phi_{\mathbf{E}}(t) = \|I_h\|_{\mathbf{E}} \quad (P(h) = t)$$

is well defined on $[0, 1]$. It is called *the fundamental function* of \mathbf{E} . The next statement follows from the results of [29], Ch. 2.

Proposition 6. *If*

$$b_{\mathbf{E}}(X) \stackrel{\text{def}}{=} \int_0^\infty \phi_{\mathbf{E}}(P\{|X| \geq x\})dx < \infty,$$

then $X \in \mathbf{E}$ and $\|X\|_{\mathbf{E}} \leq a(\mathbf{E})b_{\mathbf{E}}(X)$, where $a(\mathbf{E})$ depends on \mathbf{E} only.

2. The function of dilatation and Boyd indices

For the r.v. X and $t > 0$ let $B_t(X)$ denote the set of all r.v.s Y such that $P\{x \leq Y < y\} \leq tP\{x \leq X < y\}$ for every $x < y, xy > 0$. If $Y \in B_t(X)$, then

$$P\{|Y| \geq x\} \leq tP\{|X| \geq x\}$$

for all positive x . Proposition 2 implies that if $X \in \mathbf{E}$, then $B_t(X) \subset \mathbf{E}$ for every $t > 0$. Put

$$\gamma_{\mathbf{E}}(t) = \sup \left\{ \frac{\|X\|_{\mathbf{E}}}{\|Y\|_{\mathbf{E}}} : Y \in B_t(X), X \in \mathbf{E}, X \neq 0 \right\}.$$

We call $\gamma_{\mathbf{E}}(t)$ the *function of dilatation* of the r.i. space \mathbf{E} . It is obvious that $\gamma_{\mathbf{E}}(t)$ is non-decreasing.

Suppose that the probability space is $[0, 1]$ with Lebesgue measure and for a r.v. $X(s)$ on $[0, 1]$ putp

$$D_t X(s) = \begin{cases} X(s/t) & \text{if } s < t, \\ 0 & \text{if } t < s \leq 1. \end{cases}$$

The operator D_t is bounded in every r.i. space \mathbf{E} (see [29]). It is easy to verify that $\gamma_{\mathbf{E}}(t) = \|D_t\|_{\mathbf{E} \rightarrow \mathbf{E}}$.

Proposition 7. *The following estimates are true:*

$$\min\{1, t\} \leq \gamma_{\mathbf{E}}(t) \leq \max\{1, t\}.$$

These inequalities follow directly from the definition and the obvious relations $\gamma_{\mathbf{E}}(t) \leq 1 = \gamma_{\mathbf{E}}(1)$ if $t \leq 1$ and $\gamma_{\mathbf{E}}(t) \geq 1$ if $t > 1$.

The *Boyd indices* are defined by the formulae

$$\alpha(\mathbf{E}) = \lim_{t \rightarrow 0} \frac{\log(\gamma_{\mathbf{E}}(t))}{\log(t)}, \quad \beta(\mathbf{E}) = \lim_{t \rightarrow \infty} \frac{\log(\gamma_{\mathbf{E}}(t))}{\log(t)}.$$

These limits exist and $0 \leq \alpha(\mathbf{E}) \leq \beta(\mathbf{E}) \leq 1$ (see [29], Ch.2). We have $\alpha(L_{p,q}) = \beta(L_{p,q}) = 1/p$.

Proposition 8. *For each $\epsilon > 0$ there exist positive constants a and b such that $\gamma_{\mathbf{E}}(t) \leq t^{\alpha(\mathbf{E})-\epsilon}$ if $t < a$ and $\gamma_{\mathbf{E}}(t) \leq t^{\beta(\mathbf{E})+\epsilon}$ if $t > b$.*

Proof: See [29], Ch. 2. \square

The following assertion is well known (see [23]).

Proposition 9. *If $\alpha(\mathbf{E}) > 0$, then $\mathbf{E} \supset L_p(\Omega)$ for all $p > 1/\alpha(\mathbf{E})$. If $\beta(\mathbf{E}) < 1$, then $\mathbf{E} \subset L_q(\Omega)$ for each $q < 1/\beta(\mathbf{E})$.*

3. Independent random variables

Definition 2. *Let H be a set of indices. The r.v.s $\{X_h\}$, $h \in H$, are said to be independent, if for every finite set h_1, \dots, h_m and each $a_k < b_k$ ($1 \leq k \leq m$)*

$$P\{a_k \leq X_{h_k} \leq b_k, 1 \leq k \leq m\} = \prod_{k=1}^m P\{a_k \leq X_{h_k} \leq b_k\}.$$

Proposition 10. *Let r.v.s X and Y be independent, \mathbf{E} be a r.i. space and $X + Y \in \mathbf{E}$. Then $X \in \mathbf{E}$ and $Y \in \mathbf{E}$.*

Proof: Let α be a median of X . It means $P\{X \leq \alpha\} \geq 1/2$ and $P\{X \geq \alpha\} \geq 1/2$ (see [35]). Since X and Y are independent, then for every $x > 0$

$$\begin{aligned} & P\{|X + Y| \geq x\} \\ & \geq P\{X \geq x - \alpha, Y \geq \alpha\} + P\{X \leq -x - \alpha, Y \leq \alpha\} \\ & \geq \frac{1}{2}P\{X \geq x - \alpha\} + \frac{1}{2}P\{X \leq -x - \alpha\} \\ & \geq \frac{1}{2}P\{|X| \geq x + |\alpha|\}. \end{aligned}$$

Using Proposition 2, we get $X \in \mathbf{E}$. \square

Proposition 11. *For each r.i. space \mathbf{E} there exists a constant $C(\mathbf{E}) > 0$ such that for any independent r.v.s $X, Y \in \mathbf{E}$, $EY = 0$*

$$\|X + Y\|_{\mathbf{E}} \geq C(\mathbf{E})\|X\|_{\mathbf{E}}.$$

Proof: Suppose that this assertion is not true for some r.i. space \mathbf{E} . Then there exist r.v.s $X_n, Y_n \in \mathbf{E}$ with the following properties:

- 1) the r.v.s $\{X_k\}_{k=1}^{\infty} \cup \{Y_k\}_{k=1}^{\infty}$ are independent;
- 2) $EY_k = 0$;
- 3) $\|X_k\|_{\mathbf{E}} = 1$;
- 4) $\|X_k + Y_k\|_{\mathbf{E}} \leq 2^{-k}$.

According to 4), the series $\sum_{k=1}^{\infty} (X_k + Y_k)$ is absolutely convergent in \mathbf{E} and, therefore, in $L_1(\Omega)$. Let β be the σ -algebra generated by the sequence $\{X_k\}_{k=1}^{\infty}$. Since the operator E^β of conditional expectation is bounded in $L_1(\Omega)$, the conditions 1) and 2) imply for each $m < n$ (see [35])

$$E^\beta \left(\sum_{k=m}^n (X_k + Y_k) \right) = \sum_{k=m}^n X_k + \sum_{k=m}^n EY_k = \sum_{k=m}^n X_k.$$

Hence the series $\sum_{k=1}^{\infty} X_k$ and $\sum_{k=1}^{\infty} Y_k = \sum_{k=1}^{\infty} (X_k + Y_k) - \sum_{k=1}^{\infty} X_k$ are convergent in $L_1(\Omega)$. According to 1) and 2), the second series is convergent almost surely (see [35]). Hence $Y_k \rightarrow 0$ almost surely.

Let $h_k(\epsilon) = \{|Y_k| < \epsilon\}$, where $0 < \epsilon < 1$. We have $P(h_k(\epsilon)) \geq 1 - \epsilon$ for sufficiently large k and

$$|X_k + Y_k|I_{h_k(\epsilon)} \geq |X_k| - \epsilon I_{h_k(\epsilon)}.$$

By virtue of independence, for each $x > 0$

$$P \{ ||X_k| - \epsilon I_{h_k(\epsilon)} \geq x \} = P(h_k(\epsilon))P \{ ||X_k| - \epsilon \geq x \}.$$

From here and Proposition 2

$$|(|X_k| - \epsilon)I_{h_k(\epsilon)}|_{\mathbf{E}} \geq P(h_k(\epsilon)) ||X_k| - \epsilon|_{\mathbf{E}}.$$

Hence for sufficiently large k

$$\begin{aligned} \|X_k + Y_k\|_{\mathbf{E}} &\geq \| (X_k + Y_k) I_{h_k(\epsilon)} \|_{\mathbf{E}} \geq \| (|X_k| - \epsilon) I_{h_k(\epsilon)} \|_{\mathbf{E}} \\ &\geq P(h_k(\epsilon)) ||X_k| - \epsilon|_{\mathbf{E}} \geq (1 - \epsilon) (\|X_k\|_{\mathbf{E}} - \epsilon) \\ &= (1 - \epsilon)^2. \end{aligned}$$

This contradicts 4). \square

Remark. If \mathbf{E} is separable or maximal, then $C(\mathbf{E}) = 1$. Indeed, let β be the σ -algebra generated by X . Proposition 4 implies that the operator E^β is bounded in \mathbf{E} and has norm equal to 1. Hence

$$\|X + Y\|_{\mathbf{E}} \geq \|E^\beta(X + Y)\|_{\mathbf{E}} = \|X + EY\|_{\mathbf{E}} = \|X\|_{\mathbf{E}}.$$

For the spaces $L_p(\Omega)$ ($1 \leq p \leq \infty$) this statement is well known (see [35]).

4. Probability inequalities

Here we present some well known estimates which will be used in the following chapters.

1. Paley and Zygmund's inequality [26]. Let $\{X_k\}_{k=1}^\infty$ be independent r.v.s such that $EX_k^4 < \infty$ and $EX_k = 0$. Suppose

$$D = \sup \left\{ \frac{EX_k^4}{(EX_k^2)^2} < \infty \right\}.$$

Then for every $a_k \in \mathbf{R}$ and $n \in \mathbf{N}$

$$P \left\{ \left| \sum_{k=1}^n a_k X_k \right| \geq \lambda \left(\sum_{k=1}^n a_k^2 \right)^{1/2} \right\} \geq \eta,$$

where $\eta = (1 - \lambda^2) \min\{1/3, 1/D\}$.

2. Bernstein's inequality [43]. Let $\{X_k\}_{k=1}^\infty$ be independent r.v.s, $|X_k| \leq C$ and $EX_k = 0$. Then

$$P \left\{ \left| \sum_{k=1}^n a_k X_k \right| \geq 2Cx \left(\sum_{k=1}^n a_k^2 \right)^{1/2} \right\} \leq \exp(-x^2)$$

for all $x > 0$, $a_k \in \mathbf{R}$ and $n \in \mathbf{N}$.

3. Prokhorov's 'arcsinh' inequality [46]. Let $\{X_k\}_{k=1}^\infty$ be independent r.v.s, $|X_k| \leq C$ and $EX_k = 0$. Let $\sigma^2 = \sum_{k=1}^n EX_k^2$. Then for every $x > 0$

$$P \left\{ \left| \sum_{k=1}^n X_k \right| \geq x \right\} \leq 2 \exp \left(-\frac{x}{2C} \operatorname{arcsinh} \frac{Cx}{\sigma^2} \right).$$

A r.v. X is said to be *symmetric* if $P\{X \geq x\} = P\{X \leq -x\}$ for all $x > 0$. If F is a probability distribution on \mathbf{R} , then $\Pi(F)$ denotes the corresponding Poisson distribution, i.e. the distribution with the characteristic function

$$g(t) = \exp \left(\int_{-\infty}^{\infty} (e^{itx} - 1) dF(x) \right). \tag{4}$$

Let us recall we write $X \in \mathcal{L}(F)$ if F is the distribution of the r.v. X .

4. Prokhorov's inequality [45]. Let $\{X_k\}_{k=1}^n$ be independent symmetric r.v.s with distributions F_k and let the r.v.s $Y_k \in \mathcal{L}(\Pi(F_k))$ be independent. Then for every $x > 0$

$$P \left\{ \left| \sum_{k=1}^n X_k \right| \geq x \right\} \leq 8P \left\{ \left| \sum_{k=1}^n Y_k \right| \geq \frac{x}{2} \right\}.$$

5. Kwapien and Richlik's inequality [53]. Let $\{X_k\}_{k=1}^\infty$ and $\{Y_k\}_{k=1}^\infty$ be sequences of independent symmetric r.v.s such that $P\{|X_k| \geq x\} \leq AP\{B|Y_k| \geq x\}$ for all $x > 0$ and $k \in \mathbf{N}$. Then for every $a_k \in \mathbf{R}$, $x > 0$ and $n \in \mathbf{N}$

$$P \left\{ \left| \sum_{k=1}^n a_k X_k \right| \geq x \right\} \leq 2AP \left\{ AB \left| \sum_{k=1}^n a_k Y_k \right| \geq x \right\}.$$

5. Disjoint random variables

We say the r.v.s X and Y are *disjoint*, if $XY = 0$.

Definition 3. [33]. A r.i. space \mathbf{E} is said satisfy the upper, respectively the lower, r -estimate if there exists a positive constant A such that for each collection of mutually disjoint r.v.s $\{X_k\}_{k=1}^n \subset \mathbf{E}$

$$\left\| \sum_{k=1}^n X_k \right\|_{\mathbf{E}} \leq A \left(\sum_{k=1}^n \|X_k\|_{\mathbf{E}}^r \right)^{1/r},$$

respectively

$$\left\| \sum_{k=1}^n X_k \right\|_{\mathbf{E}} \geq A \left(\sum_{k=1}^n \|X_k\|_{\mathbf{E}}^r \right)^{1/r}.$$

The following assertions are known. For completeness the proofs are given.

Proposition 12. Let $1 \leq p < \infty$, $1 \leq q \leq \infty$ and $r = \min\{p, q\}$. The Lorentz space $L_{p,q}(\Omega)$ satisfies the upper r -estimate and does not satisfy the upper s -estimate for $s > r$.

Proof: Let $\{X_k\}_{k=1}^n \subset \mathbf{E}$ be mutually disjoint r.v.s. Then for every $x > 0$

$$P \left\{ \left| \sum_{k=1}^n X_k \right| \geq x \right\} = \sum_{k=1}^n P\{|X_k| \geq x\}.$$

Suppose $q = \infty$. According to (2)

$$\begin{aligned} \left(\left\| \sum_{k=1}^n X_k \right\|_{p,\infty}^* \right)^p &= \sup \left\{ x^p P \left\{ \left| \sum_{k=1}^n X_k \right| \geq x \right\} : x > 0 \right\} \\ &= \sup \left\{ x^p \sum_{k=1}^n P\{|X_k| \geq x\} : x > 0 \right\} \\ &\leq \sum_{k=1}^n (\|X_k\|_{p,\infty}^*)^p. \end{aligned}$$

Let $r = p \leq q < \infty$. From the formula (1)

$$\left(\left\| \sum_{k=1}^n X_k \right\|_{p,q}^* \right)^p = \left(\int_0^\infty \left(\sum_{k=1}^n P\{|X_k| \geq x\} \right)^{q/p} dx^q \right)^{p/q}.$$

Applying Minkowski's inequality with the exponent $\alpha = q/p$, we get the needed estimate.

Let's turn to the case $r = q < p$. We have

$$\left(\left\| \sum_{k=1}^n X_k \right\|_{p,q}^* \right)^q = \int_0^\infty \left(\sum_{k=1}^n P\{|X_k| \geq x\} \right)^{q/p} dx^q.$$

Since $|a + b|^t \leq |a|^t + |b|^t$ for $0 < t < 1$, then

$$\left(\sum_{k=1}^n P\{|X_k| \geq x\} \right)^{q/p} \leq \sum_{k=1}^n (P\{|X_k| \geq x\})^{q/p}.$$

From here the desired estimate follows.

Let $s > r$. If a r.i. space \mathbf{E} satisfies the upper s -estimate, then $\beta(\mathbf{E}) < 1/s$ (see [33]). Since $\beta(L_{p,q}) = 1/p$, then, if $r = p$, the space $L_{p,q}(\Omega)$ does not satisfy the upper s -estimate.

Suppose $r = q$. Then, as shown in [40], there exist mutually disjoint normed r.v.s $\{X_k\}_{k=1}^\infty \subset L_{p,q}(\Omega)$ such that for all $n \in \mathbf{N}$

$$\left\| \sum_{k=1}^n X_k \right\|_{p,q}^* \geq n,$$

where $C > 0$ is a constant. So, $L_{p,q}(\Omega)$ does not satisfy the upper s -estimate for $s > r = q$ and the proposition is proved. \square

Now we consider Orlicz spaces. Let $N_i(x)$ ($i = 1, 2$) be even convex functions on \mathbf{R} . We say that these functions are *equivalent* if there exist positive constants A, B and C such that $N_1(Ax) \leq N_2(x) \leq N_1(Bx)$ for $x > C$. It is shown in [28] that $L_{N_1}(\Omega) = L_{N_2}(\Omega)$ if and only if the functions $N_1(x)$ and $N_2(x)$ are equivalent.

Proposition 13. *Let a convex even function $N(x)$, $N(0) = 0$, be equivalent to $U(x) = |x|^r \Psi(x)$, where $r > 1$ and $\Psi(x)$ is convex and increasing on $(0, \infty)$. Then the space $L_N(\Omega)$ satisfies the upper r -estimate.*

Proof.: Without loss of generality $N(x) = |x|^r \Psi(x)$. Suppose $\{X_k\}_{k=1}^n \subset L_N(\Omega)$ are mutually disjoint random variables and $EN(\lambda_k^{-1} X_k) \leq 1$ for some real λ_k ($1 \leq k \leq n$). Put $t_k = \lambda_k (\sum_{j=1}^n \lambda_j^r)^{-1/r}$. Then

$$EN \left(\left(\sum_{j=1}^n \lambda_j^r \right)^{-1/r} \sum_{k=1}^n X_k \right) = \sum_{k=1}^n EN(t_k \lambda_k^{-1} X_k).$$

Since the function $\Psi(x)$ is even and increasing on $(0, \infty)$ we have $\Psi(tx) \leq \Psi(x)$ for $0 < t < 1$ and every $x \in \mathbf{R}$. Therefore $EN(tX) = E(|tX|^r \Psi(tX)) \leq |t|^r E(|X|^r \Psi(X)) = |t|^r EN(X)$. From here

$$EN \left(\left(\sum_{j=1}^n \lambda_j^r \right)^{-1/r} \sum_{k=1}^n X_k \right) \leq \sum_{k=1}^n t_k^r EN(\lambda_k^{-1} X_k) \leq \sum_{k=1}^n t_k^r = 1.$$

Using (3), we get

$$\left\| \sum_{k=1}^n X_k \right\|_{L_N} \leq \left(\sum_{k=1}^n t_k^r \right)^{1/r}.$$

Applying (3) once more, we obtain the desired bound. \square

6. The Kruglov property

In this section we investigate relations between independent r.v.s and mutually disjoint ones. A similar problem has been studied by Carothers and Dilworth [14], [15] and Johnson and Schechtman [24]. Our approach is different from theirs. It is based on the so-called Kruglov property of a r.i. space.

Kruglov proved the following statement [30].

Kruglov’s Theorem. *Let $\Phi(x)$ be a non-negative and continuous function on \mathbf{R} satisfying one of the following conditions:*

$$\Phi(x + y) \leq B\Phi(x)\Phi(y), \quad \Phi(x + y) \leq B(\Phi(x) + \Phi(y)),$$

where B is a constant. Suppose $X \in \mathcal{L}(F)$ and $Y \in \mathcal{L}(\Pi(F))$. Then the conditions $E\Phi(X) < \infty$ and $E\Phi(Y) < \infty$ are equivalent.

The question arises about the conditions under which the result of this type is true for r.i. spaces.